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Fine particulate air pollution, nitrogen dioxide, and systemic autoimmune rheumatic disease in Calgary, Alberta

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ABSTRACT

Objective: To estimate the association between fine particulate (PM_{2.5}) and nitrogen dioxide (NO₂) pollution and systemic autoimmune rheumatic diseases (SARDs).

Methods: Associations between ambient air pollution (PM_{2.5} and NO₂) and SARDs were assessed using land-use regression models for Calgary, Alberta and administrative health data (1993–2007). SARD case definitions were based on ≥ 2 physician claims, or ≥ 1 rheumatology billing code; or ≥ 1 hospitalization code (for systemic lupus, Sjogren's Syndrome, scleroderma, polymyositis, dermatomyositis, or undifferentiated connective tissue disease). Bayesian hierarchical latent class regression models estimated the probability that each resident was a SARD case, based on these case definitions. The sum of individual level probabilities provided the estimated number of cases in each area. The latent class model included terms for age, sex, and an interaction term between age and sex. Bayesian logistic regression models were used to generate adjusted odds ratios (OR) for NO₂ and PM_{2.5} pollutant models, adjusting for neighbourhood income, age, sex, and an interaction between age and sex. We also examined models stratified for First-Nations (FN) and non-FN subgroups.

Results: Residents that were female and/or aged > 45 had a greater probability of being a SARD case, with the highest OR estimates for older females. Independently, the odds of being a SARDs case increased with PM_{2.5} levels, but the results were inconclusive for NO₂. The results stratified by FN and non-FN groups were not distinctly different.

Conclusion: In this urban Canadian sample, adjusting for demographics, exposure to PM_{2.5} was associated with an increased risk of SARDs. The results for NO₂ were inconclusive.

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1. Introduction

Systemic autoimmune rheumatic diseases (SARDs) are complex chronic inflammatory disorders (such as systemic lupus, SLE) characterized by multi-system inflammation (Callaghan et al.,

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2007; Bernatsky et al., 2009; Aghdassi et al., 2011) and high disease burden (Panopalis et al., 2007). Although there is some evidence that SARDs may be triggered by environmental factors (Cooper et al., 1999; Cooper and Stroehla, 2003; Ballestar, 2010; Shapira et al., 2010) there remain considerable knowledge gaps regarding mechanisms through which the environment may trigger these autoimmune diseases. Fine particulate matter (PM_{2.5}) enters the body through airways and can trigger a systemic inflammatory response (van Eeden et al., 2001). This mechanism may be important in driving some of air pollution's effects in cardiovascular disease (van Eeden et al., 2012) and diabetes (Anderson et al., 2012; Chen et al., 2013). However, other components of air pollution, such as nitrogen dioxide (NO₂) are also pro-inflammatory and may also be of interest, and NO₂ has been linked to adverse health outcomes, including asthma (Cai et al., 2014), cardiac disease (Zhao et al., 2014), and diabetes (Eze et al., 2014). A very few studies have examined the association between air pollution and SARDs. We have previously estimated levels of PM_{2.5} across Alberta, using satellite imagery (Bernatsky et al., 2014). However, satellite remote sensing estimates air pollution concentrations over larger geographic areas (10 km grids) increasing the possibility of exposure misclassification. The goal of this study was to examine associations between air pollution and SARDs at a finer spatial scale. Our objective was also to estimate the degree to which a diagnosis of SARDs was associated not only with ambient PM_{2.5} levels, but also NO₂ levels. The current analyses focused on Calgary, the largest city in Alberta (almost 1.3 million residents, representing almost 40% of the province's population).

2. Methods

We examined associations between local-scale ambient NO₂ and PM_{2.5} in Calgary, Alberta using land use regression (LUR) models and administrative health data. The LUR modelling approach has been widely used for research estimating individual exposure to ambient air pollution (Johnson et al., 2013). Land-use regression models are built by including predictor variables, such as traffic, topography, and other geographic variables, in multivariable regression using monitored pollution levels as the outcome. Subsequently, levels of pollution may then be estimated for any point, using the parameter estimates derived from the regression model. This method has been identified as a preferred approach to estimating small area variations in air pollution effectively, when household level monitoring data are not available (Ryan and LeMasters, 2007; Health Effects Institute, 2010).

We used comprehensive (physician billing and hospitalization) administrative provincial health data from Alberta (1993–2007) and focussed our analyses on Calgary, since LUR estimates for PM_{2.5} and NO₂ were available for this region of Alberta, only. In Canada, each province maintains linkable databases on all residents who are recipients of a comprehensive health plan which covers physician visits and hospitalizations. In the Alberta health administrative databases, information on First Nations (FN) status is also available, at the individual level. The FN variable is defined by whether or not an individual's health premiums are paid by the First Nations and Inuit Health Branch (Health Canada) at any time point since 1994, thus indicating Treaty Status as per the Indian Act (Alberta Health and Wellness, 2004).

The databases record one physician billing code per visit (in Alberta, up to three diagnosis codes are allowed, but use of more than one code is infrequent), and all hospitalizations (with up to 25 diagnostic codes per hospitalization in Alberta). The SARD case definition was based on three algorithms: two or more physician billing claims with the International Classification of Diseases (ICD)-9 code 710 (within 2 years but at least 2 months apart); at

least one such billing code by a rheumatologist; or at least one hospitalization with an ICD-10 diagnostic code corresponding to a SARD diagnosis (M32.1, M32.8–32.9, M33–M34, M35.0, M35.8–35.9, M36.0). These diagnostic codes include SLE, Sjogren's Syndrome, scleroderma, polymyositis, dermatomyositis, and undifferentiated connective tissue disease. Within the ICD-9 classification system, in provinces where billing data are only limited to three digits, it is impossible to differentiate between these conditions (SLE, Sjogren's Syndrome, scleroderma, polymyositis, dermatomyositis, and undifferentiated connective tissue disease). Rheumatoid arthritis (and systemic vasculitis, by the way) fall under separate ICD-9 categories (714.x and 446.x respectively) and we were unable to include these groups in the current study.

Residents of Calgary could meet one or more of the three SARDs case definitions. Bayesian hierarchical latent class regression models (which do not assume any case definition to be a perfect gold standard) (Bernatsky et al., 2005) were used to estimate the probability that any given Calgary resident was a SARD case, based on their results relative to the three case definitions. For example, a resident who never fulfilled any of the case definitions would have an estimated probability of being a SARD that was much smaller than someone who fulfilled all three case definitions, the former near zero (but not exactly zero) and the latter nearer to 1 (but not exactly 1). Residents with one or two case definitions would have estimated probabilities between these extremes. The latent class approach to case definition within administrative data (as opposed to assuming chart review or anything else as a perfect gold standard), where individuals are not assigned a case status per se, but rather are provided with an estimated probability of being a case, has been shown to be useful across several diseases and jurisdictions (Prosser et al., 2008).

The individual level probabilities estimated were summed to estimate the total number of SARD cases according to groups characterized by age (dichotomized to less than or equal to 45 years of age, or older than 45), sex, and dissemination area, DA (the smallest standard geographic area for which all Canadian census data are disseminated, generally representing between 400 and 700 persons). The DA associated with the residence of each Calgary subject was obtained through overlay of Calgary postal code centroids on a geographic layer of the Calgary dissemination areas. To obtain SARD prevalence rates for each DA, for each age and sex group, we divided the estimated number of SARD cases by the appropriate regional Calgary population figures obtained from the Canadian Census 2006.

We then used the case definition results as outcome data in a logistic regression model that estimated odds ratio (OR), estimates for sex, age group, and pollution levels (explained below), across all dissemination areas in Calgary. The model also included an interaction term between age and sex. Around our point estimates of the odds ratios (ORs) generated from the model are Bayesian 95% credible intervals (CrI). Regional income information was obtained from the Canadian Census 2006 for Alberta.

Long-term exposure to ambient NO₂ and PM_{2.5} was estimated using the LUR model results developed for Calgary dissemination block areas. The LUR is based on data from two-week summer and winter air monitoring campaigns (held in August 2010 and January 2011). NO₂ LUR model data were based on data from 46 sites in summer and 47 sites in winter, while PM_{2.5} models were based on data from 25 sites in summer and 29 sites in winter. NO₂ was modelled using a geographically weighted land-use regression model to address spatial non-stationarity and autocorrelation (Beratazzon et al., 2011; Elikan et al., 2011). PM_{2.5} was modelled using traditional linear regression LUR methods, because the PM_{2.5} spatial non-stationarity and autocorrelation do not have a lot of impact on PM_{2.5} estimates. Some examples of important predictors of NO₂ and PM_{2.5} in the LUR models included traffic and

industrial sources (based on data from the National Pollutant Release Inventory, as well as road network and land-use information provided by the City of Calgary and Desktop Mapping Technologies Incorporated) and wind (meteorological data was obtained from three Calgary weather stations operated by Environment Canada). Mean annual ambient PM_{2.5} and NO₂ levels were calculated as the average of the summer and winter LUR model estimates for each pollutant.

For our analyses, which occurred at the DA level, the dissemination block-level LUR pollution estimates were aggregated within their larger dissemination area (weighted according to the block population within each DA, to best represent the exposure of the population within each DA); thus all residents within a given DA were assigned the same pollution exposure variables.

Since the risk of SARDs may be different (that is, sometimes higher) in First-Nations (FN) versus non-FN residents (Barnabe et al., 2012), we first ran models with all Calgary subjects, then we ran models where FN and non-FN residents were stratified.

3. Results

In the administrative data for the whole of Calgary, the average percentage of females was 50.2% (with the interquartile range across DAs being 48.6–51.7%), and the percentage of older individuals (> 45 years), across DAs, was 36.7% (interquartile range 29.6–43.9%). The average PM_{2.5} levels across DAs was 6.38 µg/m³ (median 6.29, interquartile range 6.03, 6.64) and the average of NO₂ across DAs was 16.30 ppb (median 16.33, interquartile range 14.44, 18.29). In Calgary, the estimated SARD prevalence, based on our methods, was about 3–4 cases per 10,000 persons.

Our hierarchical latent class logistic regression model accounted concurrently for age and sex (and an interaction term between them). Table 1 shows the results of models with all Calgary subjects, as well as models where FN and non-FN residents were stratified. We found strong evidence for effect modification, such that the OR for older females was higher by an order of magnitude compared to simply being older or being female. The effects were not distinctly different in the FN versus non-FN

Table 1

Odds ratio (OR) estimates for systemic autoimmune rheumatic disease in Calgary, for PM_{2.5}^a air pollution, adjusted for demographics (older age, sex, income) and stratified by First Nations status.

| All Calgary residents | OR | 95% | CrI ^b |
|--|------|------|------------------|
| Female sex | 8.14 | 6.15 | 10.3 |
| Mean income ^{***} | 1.00 | 0.99 | 1.00 |
| Age > 45 | 4.78 | 3.67 | 5.89 |
| Interaction term between age and sex | 24.6 | 19.3 | 31.4 |
| PM _{2.5} | 1.10 | 1.01 | 1.22 |
| First Nation residents only | | | |
| Female sex | 7.31 | 2.73 | 22.6 |
| Mean income | 1.01 | 0.88 | 1.11 |
| Age > 45 | 2.66 | 0.84 | 9.24 |
| Interaction term between age and sex | 51.0 | 17.0 | 180.4 |
| PM _{2.5} | 1.40 | 0.72 | 2.60 |
| Non-First Nation Calgary residents only | | | |
| Female sex | 7.84 | 6.06 | 10.9 |
| Mean income | 1.00 | 0.99 | 1.00 |
| Age > 45 | 4.88 | 3.63 | 6.49 |
| Interaction term between age and sex | 25.4 | 19.6 | 35.6 |
| PM _{2.5} | 1.04 | 0.94 | 1.15 |

^{***} Mean income values for each Calgary dissemination area were obtained from Canadian Census data.

^a PM_{2.5}=fine particulate air pollution, in µg/m³.

^b Bayesian credible intervals.

Table 2

Odds ratio (OR) estimates for systemic autoimmune rheumatic disease in Calgary, for nitrogen dioxide (NO₂^a) levels, adjusted for demographics (older age, sex, income) and stratified by First Nations status.

| All Calgary residents | OR | 95% | CrI ^b |
|--|------|------|------------------|
| Female Sex | 7.54 | 5.83 | 9.57 |
| Mean Income ^{***} | 0.99 | 0.98 | 1.00 |
| Age > 45 | 5.14 | 3.81 | 6.40 |
| Interaction term between age and sex | 26.6 | 20.6 | 33.3 |
| NO ₂ | 1.00 | 0.98 | 1.02 |
| First Nation residents only | | | |
| Female sex | 7.43 | 2.72 | 24.4 |
| Mean income | 1.00 | 0.87 | 1.10 |
| Age > 45 | 2.58 | 0.71 | 10.5 |
| Interaction term between age and sex | 52.3 | 17.5 | 170.8 |
| NO ₂ | 1.03 | 0.89 | 1.20 |
| Non-First Nation Calgary residents only | | | |
| Female sex | 7.80 | 6.26 | 10.02 |
| Mean income | 0.99 | 0.98 | 1.00 |
| Age > 45 | 4.77 | 3.85 | 6.44 |
| Interaction term between age and sex | 25.2 | 20.3 | 32.7 |
| NO ₂ | 0.99 | 0.98 | 1.01 |

^{***} Mean income values for each Calgary dissemination area were obtained from Canadian Census data.

^a NO₂ level estimates were expressed in parts per billion.

^b Bayesian credible intervals.

populations, but our precision was limited given that FN Calgary residents represent a relatively small number.

Independently of these effects, the odds of being a SARDs case appeared to be potentially associated with PM_{2.5} levels. Results for NO₂ were inconclusive based on the large credible intervals (Table 2).

4. Discussion

Our findings of higher rates of SARDs diagnosis for certain demographic factors, particularly female sex and older age, is consistent with both our prior work, and the known epidemiology of these diseases (Broten et al., 2014). The evidence for a potential association of SARDs with ambient PM_{2.5} is relatively novel; a very few studies have examined the association of particulate matter and the onset of autoimmune rheumatic diseases. In one prior study, PM_{2.5} levels were associated with a 60% increased risk of juvenile idiopathic arthritis in young children (Zeft et al., 2009) and in a 2009 analyses of the Nurses' Health Study by Hart et al. (2009) suggested that pollution emissions from road traffic may be an environmental risk factor for rheumatoid arthritis, RA. However, Hart et al. have authored two recent additional studies with potentially conflicting results. In a Swedish cohort, they found that exposure to NO₂, but not particulate matter (based on PM₁₀), was associated with RA risk (Hart et al., 2013). However, in another follow up study of the Nurses' Health Study (Hart et al., 2013), they failed to detect associations with either NO₂ or PM_{2.5}. Though not included in our current study, both JIA and RA are similar to the SARDs we studied, in terms of being rheumatic disease characterized by autoimmunity.

Our own team has shown the association between road-traffic density and SARD onset in Montreal (Labrecque et al., 2010) as well as suggested links between PM_{2.5} levels and SLE activity in our Montreal SLE cohort (Bernatsky et al., 2011). Very recently proinflammatory effects of particulate matter (diesel exhaust nanoparticles) were demonstrated in an in-vitro study of scleroderma skin cells (Mastrofrancesco et al., 2014), suggesting a possible mechanism for PM-mediated effects. In addition, we have recently shown the association between PM_{2.5} and SARDs

occurrence throughout the province of Alberta, based on satellite-derived estimates of ambient PM_{2.5} (Bernatsky et al., 2014). The OR for the fourth versus the first quartile of PM_{2.5} exposure in Alberta, where the effect of PM_{2.5} was non-linear, was 1.13 (95% CrI 1.01, 1.25) per 1 µg/m³. In the province of Quebec, modelling PM_{2.5} as a linear variable, we estimated an OR for SARDs of 1.055 (95% CrI 1.051, 1.059) per increase of 1 µg/m³.

Those earlier analyses of the effects of pollution on SARDs in Alberta were limited by the spatial scale and pollutants estimated by satellite remote sensing data. The analyses only assessed PM_{2.5}; although satellite remote sensing estimates are available for NO₂, they are not considered as accurate as the estimates for PM_{2.5}. Moreover, LUR estimates for air pollution provide more spatially refined, local scale estimates than other methods. In the earlier analyses based on satellite remote sensing, PM_{2.5} levels were estimated at a much larger geographic scale (i.e., 10 km × 10 km). The trade-off is that the LUR based analyses were limited to the city of Calgary, versus using province-wide data. Though use of LUR to assign exposure is considered one of the best available methods, novel approaches combining LUR estimates with satellite data are possible, and we will explore this in the future.

In our current analyses, we were able to demonstrate a potential link between PM_{2.5} and SARDs, but we were unable to draw conclusions regarding NO₂ and SARDs. It should be kept in mind that the original LUR estimates for PM_{2.5} and NO₂ were produced to reflect averages over the dissemination block, but our analyses were performed based on average concentrations at the (larger) dissemination area-level. NO₂ is more spatially heterogeneous compared with PM_{2.5}, which means that as size of the geographic unit of analysis increases, there may be increasing non-differential misclassification of NO₂. This effect would be less pronounced for PM_{2.5}, which is more spatially homogeneous. This could have led to a 'bias towards the null' in our NO₂ analyses. We note that the goodness-of-fit of the LUR model for NO₂ was actually better than the model for PM_{2.5}, although some of the standard error terms for predictors in the NO₂ LUR model were a bit larger than for PM_{2.5} model predictors (Johnson et al., 2013).

Another potential issue is temporality, because LUR estimates static concentrations at a single time period, our analyses are based on the assumption that the geographic variation of NO₂ and PM_{2.5} estimates was relatively constant for the 1993–2007 study period. Both NO₂ and PM_{2.5} levels may have decreased or increased locally with the closure or opening of new industrial sources and other land-use changes such as expanding road networks and residential populations. However, there are currently no current fine scale pollution models available for estimating historic concentrations. Community-level trends suggest that NO₂ levels have decreased over time in most Canadian cities, while PM_{2.5} levels have remained relatively constant since the late 1990s, with only a slight downward trend noted in major Canadian cities (Wood et al., 2012).

We used latent class regression models (which do not assume any case definition to be a perfect gold standard) to estimate the probability that any given Calgary resident was a SARD case, based on their results relative to the three case definitions. In the latent class approach to case definition, individuals are not assigned a case status per se, but rather are provided with an estimated probability of being a case, which been shown to be useful across several diseases and jurisdictions (Prosser et al., 2008). Our Bayesian latent class model in fact uses the information from the three case definitions to estimate, and adjust for, the imperfect nature of all case definitions, in terms of both sensitivity and specificity. A limitation of using the latent class approach, which does not identify cases per se, is that we could not with our methods distinguish between incident and prevalent SARD cases. Thus the temporal relationship between the pollutants and SARDs

incidence cannot be specifically assessed. Future analyses in incident-only SARD cases, using other approaches, could be done to examine whether the effects we saw were the same in incident versus prevalent SARD cases. Non-differential misclassification of the exposure could indeed occur if persons who had SARDs diagnoses had the same migration rate as those without SARDs diagnoses, and this would have biased the results towards the null. Differential migration on the other hand, could cause biased results, if for example greater SARDs prevalence in urban areas (where PM_{2.5} levels may be higher) is due to migration of persons from rural to urban areas after their SARDs diagnoses. Though we cannot completely rule out this potential bias, Labrecque (2013) studied migration rates in SLE and did not find rates were greater than control populations.

Residual confounding by variables incompletely captured in the current analyses, such as race/ethnicity or potentially socioeconomic status, remain as potential limitations. Though the effects were not distinctly different in the FN versus non-FN populations, our precision was limited given that FN Calgary residents represent a relatively small number.

Future studies should address whether air pollutants from other sources are also associated with SARDs. Evaluating even smaller particulate matter generated by road traffic (especially submicron particles, PM₁) would help to further characterize associations between particulate matter and SARDs. Estimates of SO₂ and PM₁ exposures are not widely available; sulphur dioxide (SO₂) from industrial emissions would also be of interest. LUR model estimates for PM₁ and SO₂ are currently being developed in Calgary, so that future analyses can examine associations between these exposures of interest regarding SARDs prevalence. Also of potential interest would be the consideration of the effects of noise, as well as air pollution, on rheumatic diseases (De Roos et al., 2014). Importantly, environmental exposure to inhaled toxic substances may be able to induce citrullination in lung cells prior to any detectable onset of inflammatory responses, suggesting that this may be an important pathway linking environmental triggers to rheumatic disease risk (Valesini et al., 2015; Farhat et al., 2011).

In conclusion, after adjusting for demographics (regional income, age and sex), our new analyses in Calgary (using LUR, a different approach to exposure estimation) replicated our earlier findings (using satellite data) suggesting that PM_{2.5} levels are associated with SARDs in Alberta. It should be kept in mind that this is an association, and that no causal inference can be made. We did not see a clear association with SARDs and NO₂ in the current analyses, but there remains additional work to do in this area.

Conflict of interest

The authors declare no conflict of interest.

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