## Technical Notes

The life table program-Three series of complete life tables are prepared by the National Center for Health Statistics (NCHS) for the U.S. population-decennial, annual preliminary, and annual final. The U.S. decennial life tables are based on decennial census data and deaths for a 3 -year period around the census year. Preliminary life tables are based on a substantial sample (approximately 90 percent) of death records. Estimates of life expectancy from the preliminary series are published annually. The annual final life tables (referred to in this section as "annual life tables") are based on a complete count of all reported deaths.

Available since 1945, the annual life tables are based on deaths occurring during the calendar year and on midyear postcensal population estimates provided by the U.S. Census Bureau. From 1945 to 1996, the annual life tables were abridged life tables and were constructed by reference to a standard table (4). Beginning with 1997 mortality data, complete life tables are constructed using a new methodology (5,6). Also for 1997, life expectancy and other life table values were shown for ages 85 to 100 years for the first time as part of the annual U.S. life tables. Previously, the annual life tables were closed at age 85. Extension of the oldest age interval was implemented by NCHS for several reasons: survival in the United States is such that approximately one-third of the population survives beyond age 85 ; improvements have occurred in age reporting at older ages; and high-quality old-age mortality data are available from the Medicare program.

Geographic coverage-The geographic areas covered in life tables before 1929-31 were limited to the death-registration areas. Life tables for 1900-1902 and 1909-11 were constructed using mortality data from the 1900 death-registration States ( 10 States and the District of Columbia) and for 1919-21 from the 1920 death-registration States (34 States and the District of Columbia). The tables for 1929-31 through 1958 cover the coterminous United States. Decennial life table values for the 3-year period 1959-61 were derived from data that include both Alaska and Hawaii for each year (tables 10 and 11). Data for each year shown in table 12 include Alaska beginning in 1959 and Hawaii beginning in 1960. However, it is not believed that the inclusion of these two States materially affects life table values.

Revised life table values, 1961-89—Life table values for 1960-69, 1970-79, and 1980-89 were constructed using the U.S. decennial life tables for 1959-61, 1969-71, and 1979-81, respectively, as the standard tables. The life table values for years prior to 1989 appearing in this publication are based on revised intercensal estimates of the populations for those years. As a result, the life table values for these years may differ from the life table values for those years published in Vital Statistics of the United States for 1989 and earlier years. Life table values for 1991 and later are based on postcensal population estimates and will be recalculated when intercensal estimates become available.

New Jersey data, 1962-64-The life tables for 1962 and 1963 for the six population groups involving race do not include data from New Jersey, which omitted the item on race from its certificates of live birth, death, and fetal death in use at the beginning of 1962. The item was restored during the latter part of 1962. However, the certificate revision without this item was used for most of 1962 as well as for 1963. For computing vital rates, populations by age, race, and sex (excluding New Jersey) were estimated to obtain comparable denominators. Approximately 7 percent of the New Jersey death records for 1964 did not
contain the race designation. When the records were being electronically processed for this State, the "race not stated" deaths were proportionally allocated to white or to black.

Nonresidents-Beginning in 1970, the deaths of nonresidents of the United States have been excluded from the life table statistics.

Estimation of life table functions-For some years, it was necessary to estimate life table functions for some race-sex groups. In tables 10 and 11, figures for the black population during the periods 1949-51 and 1959-61 were estimated using figures for the nonwhite population. Life table functions were also missing in tables 10 and 11 for race-sex groups for the periods from 1900-1902 to 1939-41. Figures were missing for the following groups:

| Years | Race and sex |
| :---: | :---: |
| 1900-1902 | Total white, total black |
| 1909-11 | Total white, total black |
| 1919-21 | Total, male, female, total white, total black |
| 1929-31 | Total, male, female, total white, total black |

These figures were estimated by weighted averages using population distributions as the weights. For example, life expectancy at age 20 years for the total black population was estimated by a weighted average of black male and black female life expectancies at age 20, using as weights the population distribution by sex of the black population age 20 years.

Annual life tables were initiated in 1945 for white males, white females, all other males, and all other females. The figures in table 12 by race and sex for the following years were estimated using a procedure other than the abridged life table methodology (12).
Years
$1900-45$ . . . . . . . . . . . . . . . . . .

Race and sex
Total
Male
Female
White
White male
White female

Annual life table functions were not calculated for the black population prior to 1970. In table 12, life expectancy for the black population for years prior to 1970 is estimated using figures for the total nonwhite population.

Population bases for computing life tables-The population used for computing life table values shown in this section (furnished by the U.S. Census Bureau) represents the resident population of the United States. The age-specific populations used for computing the 2000 life table values are based on the July 1, 2000, population estimates that are consistent with the 1990 census (13). The 1990 census counts by race and age were modified. Race was modified to be consistent with the Office of Management and Budget categories and historical categories for mortality data. The modification procedures for race and age are described in a census report (14).

Detailed populations from the 2000 census were not available when this report was prepared. A comparison of summary 2000 census results and the estimates for 2000 used in this report indicates differences for some population groups. Differences between the 2000 enumerated population and the population estimates for 2000 used in this report could result in the underestimation or overestimation of life expectancy. When the necessary population estimates based on the

2000 census become available, life expectancy and other life table estimates will be recalculated and presented in an upcoming report. Meanwhile, considerable caution should be used in interpreting the statistics presented in this report.

Medicare data-Death rates at the oldest ages based on Medicare data are known to be more accurate than those based on vital statistics and census data. Consequently, $q_{x}$ values calculated for ages 85 to 99 years are based on Medicare data prepared by the Health Care Financing Administration (HCFA). Medicare data were limited to the group insured for hospital insurance as age reporting is considered best among this group $(6,10,11)$. For the 2000 life tables, 1997 Medicare data were used as 2000 data were not available in time for the preparation of this report.

## Methodology

A more detailed treatment of the methodology used to calculate these life tables is contained in a separate report (5). Calculation of the complete life table is derived from the probability of death $\left(q_{x}\right)$, which depends on the number of deaths $\left(D_{x}\right)$ and the midyear population $\left(P_{x}\right)$ for each single year of age $(x)$ observed during the calendar year of interest.

Adjustment for deaths for which age was not reported-An adjustment must be made to account for the small proportion of deaths each year for which age is not reported. The data are aggregated into 5-year age groups for those age 5 years and over and into single years of age for those under 5 years. The number of deaths in each age category is adjusted proportionally to account for those with not-stated ages. The following factor is used to make the adjustment. This factor $(F)$ is calculated for each race-sex group for which life tables are constructed.

$$
\begin{equation*}
F=\frac{D}{D^{a}} \tag{1}
\end{equation*}
$$

where $D$ is the total number of deaths and $D^{a}$ is the total number of deaths for which age is stated. $F$ is then applied by multiplying it times the number of deaths in each age group. Table I shows values for $F$ by race and sex used to adjust the 2000 mortality data.

Interpolation of $P_{x}$ and $D_{x}$-Anomalies, both random and those associated with reporting age at death, can be problematic when using vital statistics and census data by single years of age to estimate the probability of death (1). Graduation techniques are often used to eliminate these anomalies and to derive a smooth curve by age. Beer's

## Table I. Values for $F$ used to adjust for not-stated age based on 2000 mortality data

| Race and sex | Total deaths | Total deaths for which age was not stated | F |
| :---: | :---: | :---: | :---: |
| Total. | 2,403,351 | 356 | 1.00014815 |
| Male | 1,177,578 | 289 | 1.00024548 |
| Female | 1,225,773 | 67 | 1.00005466 |
| White | 2,071,287 | 275 | 1.00013279 |
| Male | 1,007,191 | 231 | 1.00022940 |
| Female . | 1,064,096 | 44 | 1.00004135 |
| Black | 285,826 | 72 | 1.00025196 |
| Male | 145,184 | 50 | 1.00034451 |
| Female . | 140,642 | 22 | 1.00015645 |

ordinary minimized fifth difference formula is used to obtain smoothed values of $P_{x}$ and $D_{x}$ (see reference 5 for details on the application of Beer's method).

Calculation of $q_{0}-q_{0}$ is calculated by using a birth cohort method employing a separation factor ( $f$ ) defined as the proportion of infant deaths in year $t$ occurring to infants born in the previous year ( $t-1$ ). $f$ can be calculated by categorizing infant deaths by date of birth. The probability of death in the first year is calculated as

$$
\begin{equation*}
q_{0}=\frac{D_{0}(1-f)}{B^{t}}+\frac{D_{0} f}{B^{t-1}} \tag{2}
\end{equation*}
$$

where $D_{0}$ is the number of infant deaths adjusted for not-reported age, and $B^{t}$ and $B^{t-1}$ are the numbers of births in years $t$ and $t-1$, respectively. Table Il shows separation factors and numbers of births by race and sex for 1999-2000.

Calculation of $q_{x}$ for ages $1-84-q_{x}$ is calculated assuming that $I_{x}$ (number of survivors at exact age $x$ in the life table population) declines linearly between $x$ and $x+1$, i.e., that deaths between exact age $x$ and $x+1$ occur on average at age $x+1 / 2$. This simplification is generally considered acceptable when age intervals are 1 year of age in length (1). Under this assumption, $I_{x}=L_{x}+1 / 2 d_{x}$ where $L_{x}$ is the average life table population at risk of dying between ages $x$ and $x+1$ and $d_{x}$ is the number of deaths occurring between age $x$ and $x+1 . q_{x}$ is then

$$
q_{x}=\frac{d_{x}}{l_{x}}=\frac{d_{x}}{L_{x}+\frac{1}{2} d_{x}}
$$

One can make the same assumption for the observed population, i.e., that the observed population aged $x$ at risk of dying at the beginning of the year $\left(N_{x}\right)$ declines linearly between ages $x$ and $x+1$. Under this assumption, $N_{x}=P_{x}+1 / 2 D_{x}$ where $P_{x}$ is the midyear population or average observed population at risk of dying between ages $x$ and $x+1$ and $D_{x}$ is the observed number of deaths occurring between ages $x$ and $x+1 . q_{x}$ is calculated as

$$
\begin{equation*}
q_{x}=\frac{D_{x}}{N_{x}}=\frac{D_{x}}{P_{x}+\frac{1}{2} D_{x}} \tag{3}
\end{equation*}
$$

For $x=1$ to $4, D_{x}$ is the observed number of deaths adjusted for not-stated age and $P_{x}$ is obtained by Beer's interpolation formula. For $x=5$ to 84 , both $D_{x}$ and $P_{x}$ are obtained by interpolation (5).

Use of Medicare data at ages 85 to 99 years-There is ample evidence that the rate of increase in $q_{x}$ declines above age 85 ( $5,11,15-17$ ). The change in $q_{x}$ for ages over 85 years can be expressed using the formula

$$
\begin{equation*}
q_{x}=q_{x-1} \cdot e^{k_{x}} \tag{4}
\end{equation*}
$$

where $k_{x}$ denotes the age-specific rate of mortality change with age $(11,16)$. Solving for $k_{x}$ gives

$$
\begin{equation*}
k_{x}=\ln \left(q_{x}\right)-\ln \left(q_{x-1}\right) \tag{5}
\end{equation*}
$$

Values for $k_{x}$ are then obtained from the Medicare data. Table III shows values for $k$ by age, race, and sex based on 1997 Medicare data. These data show clearly a declining rate of increase in $q_{x}$ above age 85. These $k_{x}$ values are then used to obtain $q_{x}$ values for ages 85 to 99 using equation 4. This method allows for flexibility in cases where the Medicare data are not available in a timely fashion. In these cases, Medicare data for the previous year can be used to calculate $k_{x}$ values. Finally, ${ }_{\infty} q_{100}$ is set equal to 1.0 since all will die at some point in this open-ended age interval. Once $q_{x}$ is obtained for each

Table II. Births in 1999 and 2000, deaths in 2000 of infants born in 1999 and 2000, and separation factors by race and sex: United States

|  | Total |  |  | White |  |  | Black |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Both sexes | Male | Female | Both sexes | Male | Female | Both sexes | Male | Female |
| Births |  |  |  |  |  |  |  |  |  |
| 1999 | 3,959,417 | 2,026,854 | 1,932,563 | 3,132,501 | 1,605,603 | 1,526,898 | 605,970 | 307,670 | 298,300 |
| 2000 | 4,058,814 | 2,076,969 | 1,981,845 | 3,194,005 | 1,636,081 | 1,557924 | 622,598 | 316,115 | 306,483 |
| Deaths in 2000 of infants born in |  |  |  |  |  |  |  |  |  |
| 1999 | 3,461 | 1,925 | 1,536 | 2,225 | 1,244 | 981 | 1,072 | 592 | 480 |
| 2000. | 24,578 | 13,797 | 10,782 | 15,921 | 8,935 | 6,986 | 7,701 | 4,311 | 3,391 |
| Separation factor (f) | 0.123 | 0.122 | 0.125 | 0.123 | 0.122 | 0.123 | 0.122 | 0.121 | 0.124 |

Table III. $\boldsymbol{k}$ values by age, race, and sex based on insured Medicare data: United States, 1997

| Age | Total |  |  | White |  |  | Black |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Both sexes | Male | Female | Both sexes | Male | Female | Both sexes | Male | Female |
| 84-85 | 0.092590 | 0.089728 | 0.103281 | 0.093742 | 0.09136 | 0.10428 | 0.071864 | 0.066047 | 0.082589 |
| 85-86 | 0.090210 | 0.087018 | 0.100251 | 0.091842 | 0.08897 | 0.10185 | 0.070794 | 0.064457 | 0.081079 |
| 86-87 | 0.087830 | 0.084308 | 0.097221 | 0.089942 | 0.08658 | 0.09942 | 0.069724 | 0.062867 | 0.079569 |
| 87-88 | 0.085450 | 0.081598 | 0.094191 | 0.088042 | 0.08419 | 0.09699 | 0.068654 | 0.061277 | 0.078059 |
| 88-89 | 0.083070 | 0.078888 | 0.091161 | 0.086142 | 0.0818 | 0.09456 | 0.067584 | 0.059687 | 0.076549 |
| 89-90 | 0.080690 | 0.076178 | 0.088131 | 0.084242 | 0.07941 | 0.09213 | 0.066514 | 0.058097 | 0.075039 |
| 90-91 | 0.078310 | 0.073468 | 0.085101 | 0.082342 | 0.07702 | 0.0897 | 0.065444 | 0.056507 | 0.073529 |
| 91-92 | 0.075930 | 0.070758 | 0.082071 | 0.080442 | 0.07463 | 0.08727 | 0.064374 | 0.054917 | 0.072019 |
| 92-93 | 0.073550 | 0.068048 | 0.079041 | 0.078542 | 0.07224 | 0.08484 | 0.063304 | 0.053327 | 0.070509 |
| 93-94 | 0.071170 | 0.065338 | 0.076011 | 0.076642 | 0.06985 | 0.08241 | 0.062234 | 0.051737 | 0.068999 |
| 94-95 | 0.068790 | 0.062628 | 0.072981 | 0.074742 | 0.06746 | 0.07998 | 0.061164 | 0.050147 | 0.067489 |
| 95-96 | 0.066410 | 0.059918 | 0.069951 | 0.072842 | 0.06507 | 0.07755 | 0.060094 | 0.048557 | 0.065979 |
| 96-97 | 0.064030 | 0.057208 | 0.066921 | 0.070942 | 0.06268 | 0.07512 | 0.059024 | 0.046967 | 0.064469 |
| 97-98 | 0.061650 | 0.054498 | 0.063891 | 0.069042 | 0.06029 | 0.07269 | 0.057954 | 0.045377 | 0.062959 |
| 98-99 | 0.059270 | 0.051788 | 0.060861 | 0.067142 | 0.0579 | 0.07026 | 0.056884 | 0.043787 | 0.061449 |

single year of age, the other life table functions may be easily calculated.

Survivor function $\left(I_{x}\right)$-The life table radix, $I_{0}$, is set at 100,000 . For ages greater than 0 , the number of survivors remaining at exact age $x$ is calculated as

$$
\begin{equation*}
I_{x}=I_{x-1}\left(1-q_{x-1}\right) \tag{6}
\end{equation*}
$$

Decrement function $\left(d_{x}\right)$-The number of deaths occurring between age $x$ and $x+1$ is calculated from the survivor function.

$$
\begin{equation*}
d_{x}=I_{x}-I_{x+1}=I_{x} q_{x} \tag{7}
\end{equation*}
$$

Note that ${ }_{\infty} d_{100}={ }_{\infty} I_{100}$ since ${ }_{\infty} q_{100}=1.0$.
Person-years lived $\left(L_{x}\right)$-Person-years lived for ages 1 to 99 years is calculated assuming that the survivor function declines linearly between age $x$ and $x+1$. This gives the formula

$$
\begin{equation*}
L_{x}=\frac{1}{2}\left(I_{x}+I_{x+1}\right)=I_{x}-\frac{1}{2} d_{x} \tag{8}
\end{equation*}
$$

For $x=0$, the separation factor $f$ is used to calculate $L_{0}$.

$$
L_{0}=f l_{0}+(1-f) I_{1}
$$

${ }_{\infty} L_{100}$ is calculated by surviving the life table cohort from age 100 using equations 4,5 , and 6 until $L_{x}$ at these ages is essentially zero (somewhere between ages 110 and 120). $q_{x}$ for these ages can be
extrapolated from the Medicare data using equation 4 . However, $k_{x}$ values must be estimated for these ages. $k_{x}$ can be modeled as a linear function of age

$$
\begin{equation*}
k_{x}=k_{85}+(x-85) s \tag{9}
\end{equation*}
$$

where $s$ is the slope of the change in $k_{x}$ by age and $k_{85}$ is calculated as $\left[\ln \left(q_{88} / q_{81}\right)\right] / 7$ in order to minimize the effects of random fluctuations $(11,17)$. $s$ can be obtained by treating equation 9 as a linear regression model. Calculated values for $s$ are shown in table IV. The predicted values for $k_{\mathrm{x}}$ are then used to calculate $q_{\mathrm{x}}$ above age 100 using equation 4 . The corresponding $L_{x}$ values for ages 100 years and over are then summed to give $L_{100}$.

Person-years lived at and above age $x\left(T_{x}\right)-T_{x}$ is calculated by summing $L_{x}$ values at and above age $x$.

$$
\begin{equation*}
T_{x}=\sum_{t=0}^{\infty} L_{x+t} \tag{10}
\end{equation*}
$$

Life expectancy at age $x\left(e_{x}\right)$-Life expectancy at exact age $x$ is calculated as

$$
\begin{equation*}
e_{x}=\frac{T_{x}}{I_{x}} \tag{11}
\end{equation*}
$$

Table IV. Slope of the change in $k$ values (s) by race and sex

| Race and sex | $s$ |
| :---: | :---: |
| Total, both sexes | -0.002379 |
| Male | -0.002710 |
| Female . | -0.003031 |
| White, both sexes . | -0.001902 |
| Male | -0.002390 |
| Female . | -0.002427 |
| Black, both sexes . | -0.001074 |
| Male | -0.001586 |
| Female . | -0.001512 |

## Abridging the complete life table

An abridged or collapsed version of the complete life table can be easily calculated in which life table functions are shown for 5-year rather than single-year age intervals. It is often desirable to summarize the life table and save space when publishing life table data by single years of age (18). The abridgement of the complete life table is simplified by an important property of three of the six life table functions. The $I_{x}, T_{x}$, and $e_{x}$ functions describe exact age $x$, i.e., the beginning of the age interval $x$ to $x+n$ ( $n$ denotes the length of the age interval-for 5 -year age intervals $n=5$ ). Life expectancy at age $20\left(e_{20}\right)$, for example, has the same value regardless of whether the
age interval is $20-21$ years or $20-25$ years. Thus, the values $I_{x}, T_{x}$, and $e_{x}$ can be extracted at 5 -year intervals from the complete life table and placed into the abridged life table (compare $I_{x}, T_{x}$, and $e_{x}$ in table V with the same functions in table 1). It is also illustrative to compare values for $e_{x}$ and $I_{x}$ in tables A and B with their corresponding values presented in tables $1-9$. The $q_{x}, d_{x}$, and $L_{x}$ functions, in contrast, describe the age interval $x$ to $x+n$. In fact, for abridged life tables, the notation for these functions is different $\left({ }_{n} q_{x},{ }_{n} d_{x}\right.$, and $\left.{ }_{n} L_{x}\right)$. Thus, ${ }_{5} q_{20}$ is the probability of dying between ages 20 and 25 years and will obviously be somewhat larger than $q_{20}$, the probability of dying between ages 20 and 21 years. Taking this into account, ${ }_{n} q_{x},{ }_{n} d_{x}$, and ${ }_{n} L_{x}$ must be recalculated in the abridged life table. It is simplest to begin with ${ }_{n} d_{x}$. The calculations are made for all but the final age interval as follows:

$$
\begin{aligned}
& { }_{n} d_{x}=I_{x}-I_{x+n} \\
& { }_{n} q_{x}=\frac{{ }_{n} d_{x}}{I_{x}} \\
& { }_{n} L_{x}=T_{x}-T_{x+n}
\end{aligned}
$$

Note that for the open-ended interval, ages 100 and over: ${ }_{\infty} d_{100}=I_{100, \infty} q_{100}=1.0$, and $_{\infty} L_{100}=T_{100}$. Table V shows each of the life table functions for the 2000 U.S. total population abridged from table 1.

Table V. Abridged life table for the total population: United States, 2000

|  | Probability of dying between ages $x$ to $x+n$ | Number surviving to age $x$ | Number dying between ages $x$ to $x+n$ | Person-years lived between ages $x$ to $x+n$ | Total number of person-years lived above age $x$ | Expectation of life at age $x$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age | ${ }_{n} q_{x}$ | $I_{x}$ | ${ }_{n} d_{x}$ | ${ }_{n} L_{x}$ | $T_{x}$ | $e_{x}$ |
| 0-1 | 0.00693 | 100,000 | 693 | 99,392 | 7,686,810 | 76.9 |
| 1-5 | 0.00131 | 99,307 | 130 | 396,916 | 7,587,418 | 76.4 |
| 5-10 | 0.00082 | 99,177 | 82 | 495,668 | 7,190,502 | 72.5 |
| 10-15 | 0.00104 | 99,095 | 103 | 495,278 | 6,694,833 | 67.6 |
| 15-20 | 0.00341 | 98,992 | 338 | 494,200 | 6,199,555 | 62.6 |
| 20-25 | 0.00479 | 98,654 | 473 | 492,113 | 5,705,355 | 57.8 |
| 25-30 | 0.00494 | 98,181 | 485 | 489,702 | 5,213,242 | 53.1 |
| 30-35 | 0.00578 | 97,696 | 565 | 487,130 | 4,723,539 | 48.3 |
| 35-40 | 0.00806 | 97,132 | 783 | 483,813 | 4,236,409 | 43.6 |
| 40-45 | 0.01182 | 96,349 | 1,139 | 479,070 | 3,752,596 | 38.9 |
| 45-50 | 0.01773 | 95,210 | 1,688 | 472,085 | 3,273,527 | 34.4 |
| 50-55 | 0.02576 | 93,522 | 2,409 | 461,940 | 2,801,442 | 30.0 |
| 55-60 | 0.03968 | 91,113 | 3,615 | 447,124 | 2,339,510 | 25.7 |
| 60-65 | 0.06133 | 87,498 | 5,366 | 424,879 | 1,892,377 | 21.6 |
| 65-70 | 0.09217 | 82,131 | 7,570 | 392,758 | 1,467,498 | 17.9 |
| 70-75 | 0.13838 | 74,561 | 10,317 | 348,168 | 1,074,739 | 14.4 |
| 75-80 | 0.20557 | 64,244 | 13,207 | 289,331 | 726,571 | 11.3 |
| 80-85 | 0.31503 | 51,037 | 16,078 | 215,947 | 437,240 | 8.6 |
| 85-90 | 0.46111 | 34,959 | 16,120 | 133,503 | 221,293 | 6.3 |
| 90-95 | 0.61506 | 18,839 | 11,587 | 62,766 | 87,790 | 4.7 |
| 95-100. | 0.75434 | 7,252 | 5,470 | 20,388 | 25,024 | 3.5 |
| 100 years and over | 1.00000 | 1,781 | 1,781 | 4,636 | 4,636 | 2.6 |


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